

## Estimating productivity of macrobenthic invertebrates from biomass and mean individual weight<sup>1</sup>

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### Abstract

The empirical relations between annual somatic production  $P$ , annual  $P/\bar{B}$ -ratio, mean annual biomass  $\bar{B}$  and mean individual weight  $\bar{W}$  in macrobenthic invertebrate populations from areas between 12° and 70° latitude and down to 300 m water depth were analysed by means of multiple linear regression of the log-transformed data. In total, 337 data sets, representing 138 species were included.

The dependence of  $P$  on  $\bar{B}$  and  $\bar{W}$  is highly significant, whereas  $P/\bar{B}$  depends only on  $\bar{W}$  and not on  $\bar{B}$ . The regression equations differ significantly (Prob. < 0.01) among the main taxonomic groups (Mollusca, Polychaeta, Crustacea).

The comparison of original data on the production of macrobenthic populations with estimates based on the empirical relations established here shows that at the community level (i.e. sum of all species) annual production can be estimated from mean annual biomass  $\bar{B}$  and mean individual weight  $\bar{W}$  of each population by means of the empirical regression functions of  $P$  on  $\bar{B}$  and  $\bar{W}$ .

### Kurzfassung

Die Schätzung der Produktivität makrobenthischer Evertebraten anhand von Biomasse und mittlerem Individualgewicht

Die empirischen Beziehungen zwischen jährlicher somatischer Produktion  $P$ , jährlicher  $P/\bar{B}$ -Rate, mittlerer Biomasse  $\bar{B}$  und mittlerem Individualgewicht  $\bar{W}$  makrobenthischer Evertebraten wurden untersucht. 337 Datensätze (138 verschiedene Arten) von Populationen aus Gebieten zwischen 12° und 70° geographischer Breite und aus Wassertiefen bis 300 m wurden mittels multipler linearer Regression der log-transformierten Werte analysiert.

Die Beziehung zwischen  $P$  sowie  $\bar{B}$  und  $\bar{W}$  ist hoch signifikant, während die  $P/\bar{B}$ -Rate nur von  $\bar{W}$  und nicht von  $\bar{B}$  abhängt. Zwischen den taxonomischen Hauptgruppen (Mollusca, Polychaeta, Crustacea) bestehen signifikante Unterschiede (Prob. < 0.01).

Der Vergleich von Originaldaten zur Produktion makrobenthischer Populationen mit Schätzungen auf der Basis der hier erstellten empirischen Beziehungen zeigt, daß die Produktion von benthischen Gemeinschaften (d.h. Summe aller Arten) anhand der mittleren Biomasse  $\bar{B}$  und des mittleren Individualgewichts  $\bar{W}$  jeder Art geschätzt werden kann.

### Resumen

Estimación de la productividad de invertebrados macrobentónicos usando biomasa y peso individual promedio

Se analizó las relaciones empíricas entre la producción somática anual  $P$ , tasa  $P/\bar{B}$  anual, biomasa media anual  $\bar{B}$  y peso medio individual  $\bar{W}$  en poblaciones de invertebrados macrobentónicos de áreas ubicadas entre 12° y 70° de latitud y hasta 300 m de profundidad. Se realizaron análisis de regresión lineal múltiple con datos log-transformados de 337 sets de datos representando a 138 especies.

$P$  depende con una alta significancia de  $\bar{B}$  y  $\bar{W}$ , mientras que  $P/\bar{B}$  depende solamente de  $\bar{W}$  y no de  $\bar{B}$ .

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$\bar{B}$ . Las ecuaciones de regresión difieren significativamente (Prob. < 0.01) entre los principales grupos taxonómicos (Mollusca, Polychaeta, Crustacea).

La comparación de los datos originales de producción de las poblaciones macrobentónicas con las estimaciones basadas en las relaciones empíricas establecidas aquí, muestra que a nivel comunitario (i.e. suma de todas las especies) la producción anual puede ser estimada a partir de la biomasa media anual  $\bar{B}$  y del peso medio individual  $\bar{W}$  de cada población por medio de funciones empíricas de regresión de  $P$  sobre  $\bar{B}$  y  $\bar{W}$ .

## Introduction

The somatic production of macrozoobenthic populations is an important component of energy flow and organic matter cycling in aquatic ecosystems, especially in shallow water areas. With respect to the exploitation of demersal fish and shellfish stocks, benthic secondary production is also of economic importance.

The assessment of production, however, is a time-consuming and expensive work, even at the level of a single population, and therefore many colleagues searched for an easier way to get a reliable estimate of production. MANN (1969) suggested that there may be a general relationship between the annual production/biomass-ratio ( $P/\bar{B}$ -ratio) and life history for most organisms. Since that time, several publications have dealt with the dependence of the  $P/\bar{B}$ -ratio on a certain aspect of the life history of an organism. WATERS (1977) showed that the annual  $P/\bar{B}$ -ratio of a population increases with an increasing number of generations per year. ZAKA (1970), ROBERTSON (1979) and WARWICK (1980) demonstrated that the  $P/\bar{B}$ -ratio decreases with the life span of the organisms. The  $P/\bar{B}$ -ratio was found to be inversely related also to weight at sexual maturity (BANSE and MOSHER 1980; HEIP et al. 1982) and to individual body mass per size class (SCHWINGHAMER et al. 1986). All of these empirical findings depend on the negative exponential relation of metabolic rate to body weight in animals.

Unfortunately, these relations between  $P/\bar{B}$ -ratio and one independent variable may not be valid for an accurate estimation of the productivity of a single population, because the confidence intervals of the log-log regression lines may cover a range of more than a hundred percent on the linear scale (see e.g. BANSE and MOSHER 1980).

The first aim of this paper is to establish improved empirical relations, which may allow a more accurate estimate of productivity from parameters which are easy to obtain. The most promising approach seems to be the use of more than one independent variable, as proposed first by HUMPHREYS (1980), who used biomass and maximum individual weight to predict assimilation and production.

The second aim of this paper is a suggestion of WARWICK (1980), who assumed that it may be possible not to estimate the production of a single population, but the production of the whole community in question by empirical relationships of this kind.

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## Methods

I investigated the relations between annual somatic production ( $P$ ), annual  $P/\bar{B}$ -ratio ( $P/\bar{B}$ ), mean annual biomass ( $\bar{B}$ ) and mean individual weight ( $\bar{W}$ ). The latter is assumed to be less affected by sampling error than maximum individual weight, which was used by HUMPHREYS (1980).

The analysis is based on 107 publications on macrobenthic productivity which were selected from the available literature until 1989. The majority of the authors used the increment or removal summation method according to CRISP (1984). Some authors calculated somatic production from growth and mortality rates or by means of the size-frequency method (HYNES and COLEMAN 1968; MENZIE 1980). In a few cases somatic

production was obtained from an energy flow analysis. If necessary, weight data were converted to ash free dry weight (AFDW) using conversion factors taken from RUMOHRE et al. (1987), BREY et al. (1988) or the original authors. Generally, mean individual weight  $\bar{W}$  had to be calculated from biomass and abundance data, i.e. they are not statistically independent. In some cases abundance had to be estimated from figures showing annual variations in numbers. In order to exclude data which may be biased due to sampling and measurement errors (very low figures of biomass, weight, production), all data sets included comply with the following biological conditions:

- Mean annual biomass  $\bar{B}$ :  $\bar{B} > 0.01$  gAFDW  $m^{-2}$
- Mean individual weight  $\bar{W}$ :  $\bar{W} > = 0.00001$  gAFDW
- Annual production  $P$ :  $P > = 0.01$  gAFDW  $m^{-2} y^{-1}$

Aquatic larvae of insects (e.g. Chironomidae, Ephemeroptera) have been excluded completely because their life history is unique among macrobenthic animals.

The general form of the relations we are interested in is:

$$P = f(\bar{B}, \bar{W}) \quad (1)$$

$$P/\bar{B} = f(\bar{B}, \bar{W}) \quad (2)$$

The aim of this paper is to find equations, which are not too complex and which describe these multiple relations best. In most cases the empirical relations between different parameters of population energetics are nonlinear and a linearization by log-transformation allows the application of a linear regression model (see e.g. BANSE and MOSHER 1980; HUMPHREYS 1979; McNEIL and LAWTON 1970; ROBERTSON 1979). In contrast to HUMPHREYS (1980), who integrated biomass and weight in one variable ( $B/W^{0.75}$ ), a multiple regression model is used here:

$$\log(P) = a + b_1 \cdot \log(\bar{B}) + b_2 \cdot \log(\bar{W}) \quad (3)$$

$$\log(P/\bar{B}) = a' + b_1' \cdot \log(\bar{B}) + b_2' \cdot \log(\bar{W}) \quad (4)$$

However, a few transformations of eq. (3) show that eq. (3) and eq. (4) are nearly equivalent:

$$\log(P) = a + (b_1 - 1) \cdot \log(\bar{B}) + b_2 \cdot \log(\bar{W}) + \log(\bar{B})$$

$$\log(P) - \log(\bar{B}) = a + (b_1 - 1) \cdot \log(\bar{B}) + b_2 \cdot \log(\bar{W})$$

$$\log(P/\bar{B}) = a + (b_1 - 1) \cdot \log(\bar{B}) + b_2 \cdot \log(\bar{W}) \quad (4')$$

At the right side, eq. (3) and (4') only differ in the coefficient of  $\log(\bar{B})$ ,  $b_1$  and  $(b_1 - 1)$ , respectively.

The data were log-transformed and first examined for significant correlations by means of the calculation of the partial correlation coefficients, which express the degree of correlation between two variables after the effect of an additional variable has been removed.

Subsequently, multiple linear regressions were calculated at the level of all data sets and at the level of main taxonomic groups. Those equations which showed significant dependences were applied to data on macrobenthic community production published by several authors; the estimates of  $P$  were compared with the original production figures.

## Results

### The data base

A total of 337 data sets of  $P$ ,  $P/\bar{B}$ ,  $\bar{B}$  and  $\bar{W}$ , referring to 138 species (Table 1), was selected. Most of these were of marine or brackish water origin (111 species, 281 data sets), only a minority represents animals inhabiting freshwater sites (27 species, 56 data sets). The data

Table 1. List of macrobenthic species included in the analysis

Mollusca	<i>Pectinaria koreni</i>
<i>Abra nitida</i>	<i>Sabellaria spinulosa</i>
<i>Abra prismatica</i>	<i>Sigalion mathildae</i>
<i>Ancylus fluviatilis</i>	<i>Spiophanes hombyx</i>
<i>Anodonta anatina</i>	<i>Spiophanes kroyeri</i>
<i>Anodonta piscinalis</i>	<i>Terebellides stroemi</i>
<i>Arctica islandica</i>	<i>Tharyx marioni</i>
<i>Bellamyia unicolor</i>	Crustacea
<i>Cardium edule</i>	<i>Ampelisca agassizi</i>
<i>Chione caricellata</i>	<i>Ampelisca abraeana</i>
<i>Chlamys islandica</i>	<i>Ampelisca brevicornis</i>
<i>Chlamys varia</i>	<i>Ampelisca tenuicornis</i>
<i>Cleopatra baltimoides</i>	<i>Ampelisca typica</i>
<i>Corbicula africana</i>	<i>Asellus aquaticus</i>
<i>Donax incarnatus</i>	<i>Calocaris macandreae</i>
<i>Donax vittatus</i>	<i>Cirolana imposita</i>
<i>Dreissena polymorpha</i>	<i>Corophium insidiosum</i>
<i>Ensis directus</i>	<i>Crangonyx racomn.</i>
<i>Ensis siliqua</i>	<i>Diastylis ratkoi</i>
<i>Gyrulus acronicus</i>	<i>Emerita analoga</i>
<i>Gyrulus deflectus</i>	<i>Ericthonius fasciatus</i>
<i>Hiattella byssifera</i>	<i>Gammarus sepiacauda</i>
<i>Hydrobia neglecta</i>	<i>Gammarus pseudolimn.</i>
<i>Hydrobia sp.</i>	<i>Gammarus pulex</i>
<i>Hydrobia ulvae</i>	<i>Hyatella azteca</i>
<i>Hydrobia ventrosa</i>	<i>Mesidoboea entomon</i>
<i>Lisarca miliaris</i>	<i>Pontoporeia affinis</i>
<i>Littorina saxatilis</i>	<i>Pontoporeia lemnorata</i>
<i>Lymnaea peregra</i>	<i>Proasellus coxalis</i>
<i>Macoma balthica</i>	<i>Unicola inermis</i>
<i>Macoma calcarata</i>	Echinodermata
<i>Macoma tuberculata</i>	<i>Amphura filiformis</i>
<i>Mercenaria mercen.</i>	<i>Mora atropes</i>
<i>Mesodesma donacium</i>	<i>Ophiotrix fragilis</i>
<i>Modiolus demissus</i>	<i>Ophiura texarata</i>
<i>Mya arenaria</i>	<i>Strongylocent. dr.</i>
<i>Mya truncata</i>	<i>Strongylocent. int.</i>
<i>Mytilus edulis</i>	Miscellaneous
<i>Nacella concinna</i>	<i>Brancheia sower.</i>
<i>Nacella annulata</i>	<i>Glossiphonia heter.</i>
<i>Nacella nitidosa</i>	<i>Helobdella stagnalis</i>
<i>Nacella paxillata</i>	<i>Herpobdella octoc.</i>
<i>Nacella paxillata</i>	<i>Herpobdella test.</i>
<i>Nacella turgida</i>	<i>Lummodilus spp.</i>
<i>Pandora gouldiana</i>	<i>Tubifex castatus</i>
<i>Pharus legumen</i>	
<i>Physa gyrina</i>	
<i>Pillucina neglecta</i>	
<i>Pisidium compressum</i>	
<i>Pisidium sp.</i>	

sources are listed in the appendix, the complete data set is available on request from the author. In Table 2, the median is given as an average value, together with the range of values for the main taxonomic groups and the total data, because the distributions of  $\bar{B}$ ,  $\bar{W}$  and  $\bar{P}$  are strongly righttailed.

Fig. 1 and Fig. 2 show the distribution of the total data with respect to geographical latitude and water depth. The bulk of the data represents populations from northern boreal intertidal and shallow subtidal sites.

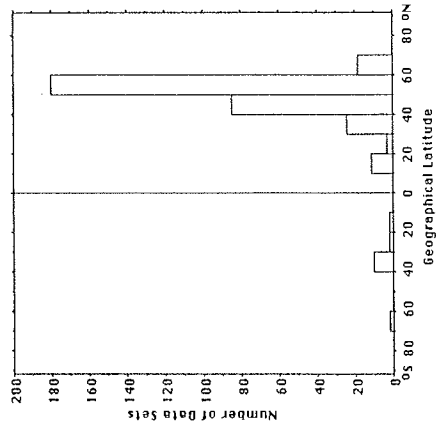


Fig. 1. Distribution of the total data with respect to geographical latitude (°N and °S).

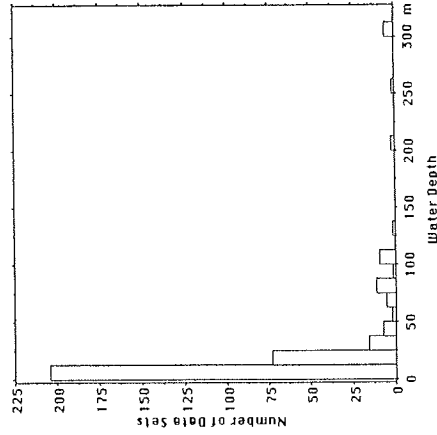


Fig. 2. Distribution of the total data with respect to water depth (m).

#### Empirical relations between $\bar{P}$ , $\bar{P}/\bar{B}$ , $\bar{B}$ and $\bar{W}$

The partial correlation coefficients based on the total data (Table 3) indicate a strong positive correlation between  $\bar{P}$  and  $\bar{B}$  ( $r = 0.909$ ) and a somewhat weaker negative correlation between  $\bar{P}$  and  $\bar{W}$  ( $r = 0.589$ ).  $\bar{P}/\bar{B}$  is not correlated with  $\bar{B}$ , but there is a negative correlation between  $\bar{P}/\bar{B}$  and  $\bar{W}$ , which is as strong as the one between  $\bar{P}$  and  $\bar{W}$ .

The corresponding multiple linear regression equations are shown in Table 3. As expected (see methods), the two equations differ only in the slope  $b_1$ . Therefore equ. (4) was not used further in the analysis.

The data were tested on significant differences among the main taxonomic groups (see

Table 2. Distribution of data

$\bar{B}$ : mean annual biomass [gAFDW m<sup>-2</sup>]  
 $W$ : mean individual weight [gAFDW]  
 $P$ : annual production [gAFDW m<sup>-2</sup> y<sup>-1</sup>]  
 $P/\bar{B}$ : annual P/ $\bar{B}$ -ratio [y<sup>-1</sup>]  
 $N_{sp}$ ,  $N_{dat}$ : Number of species, number of data sets

	$N_{sp}$	$N_{dat}$	median	range
Mollusca	$\bar{B}$	182	2.14	0.02 - 1208.00
	$W$		0.00600	0.00004 - 22.071
	$P$		1.76	0.02 - 1545.00
Polychaeta	$P/\bar{B}$	72	1.13	0.05 - 15.11
	$B$		0.85	0.02 - 20.12
	$W$		0.00200	0.00003 - 0.973
Crustacea	$P$	65	2.44	0.02 - 50.76
	$P/\bar{B}$		1.83	0.17 - 12.67
	$B$		0.43	0.02 - 52.00
Echinodermata	$W$	7	0.00100	0.00008 - 0.346
	$P$		1.02	0.08 - 217.50
	$P/\bar{B}$		4.00	0.12 - 11.29
Miscellaneous	$B$	6	7.34	0.20 - 21.00
	$W$		0.06600	0.02000 - 0.990
	$P$		6.47	0.10 - 31.50
Total	$P/\bar{B}$	7	0.87	0.48 - 1.82
	$B$	11	0.33	0.01 - 0.88
	$W$		0.00200	0.00010 - 0.007
Total	$P$	138	1.27	0.03 - 4.07
	$P/\bar{B}$		3.50	2.38 - 12.50
	$W$		1.09	0.01 - 1208.00
Total	$P$		0.00200	0.00003 - 22.071
	$P/\bar{B}$		1.77	0.02 - 1545.00
	$W$		1.82	0.05 - 15.11

Table 3. Partial correlation coefficients and multiple linear regression of P and P/ $\bar{B}$  on  $\bar{B}$  and W, total data (N = 337)

$\bar{B}$ : mean annual biomass [gAFDW m<sup>-2</sup>]  
 $W$ : mean individual weight [gAFDW]  
 $P$ : annual production [gAFDW m<sup>-2</sup> y<sup>-1</sup>]  
 $P/\bar{B}$ : annual P/ $\bar{B}$ -ratio [y<sup>-1</sup>]

Y	Partial correlation coefficients:		Probability of error	R <sup>2</sup>
	W removed	$\bar{B}$ removed		
$P$	$r_{P/\bar{B},W} = 0.909$	$r_{P,W,\bar{B}} = -0.589$		
$P/\bar{B}$	$r_{P/\bar{B},W} = 0.015$	$r_{P/\bar{B},\bar{B}} = -0.587$		
Multiple linear regression				
Model: $\log(Y) = a + b_1 \cdot \log(\bar{B}) + b_2 \cdot \log(W)$				
Coefficients				
	95 % interval		Probability of error	
$P$	a	-0.473		0.851
	b <sub>1</sub>	1.007	±0.050	<0.0001
	b <sub>2</sub>	-0.274	±0.041	<0.0001
$P/\bar{B}$	a	-0.473		0.478
	b <sub>1</sub>	0.007	±0.050	<0.7798
	b <sub>2</sub>	-0.274	±0.041	<0.0001

Table 4. Multiple linear regression of P on  $\bar{B}$  and W, main taxonomic groups

$\bar{B}$ : mean annual biomass [gAFDW m<sup>-2</sup>]  
 $W$ : mean individual weight [gAFDW]  
 $P$ : annual production [gAFDW m<sup>-2</sup> y<sup>-1</sup>]  
 $P/\bar{B}$ : annual P/ $\bar{B}$ -ratio [y<sup>-1</sup>]  
 Model:  $\log(P) = a + b_1 \cdot \log(\bar{B}) + b_2 \cdot \log(W)$

	Coefficients	95 % interval	Probability of error	R <sup>2</sup>
Mollusca	a	-0.591		0.869
	b <sub>1</sub>	1.030	±0.065	<0.0001
	b <sub>2</sub>	-0.283	±0.057	<0.0001
Polychaeta	a	-0.018		0.861
	b <sub>1</sub>	-1.022	±0.116	<0.0001
	b <sub>2</sub>	-0.116	±0.081	<0.0057
Crustacea	a	-0.614		0.913
	b <sub>1</sub>	1.022	±0.081	<0.0001
	b <sub>2</sub>	-0.360	±0.098	<0.0001
Total	a	1.007		0.851
	b <sub>1</sub>	1.007	±0.050	<0.0001
	b <sub>2</sub>	-0.274	±0.041	<0.0001

Table 2) by means of the inclusion of these groups as additional dummy variables in the regression model. Mollusca, Polychaeta and Crustacea have significantly different (Prob. < 0.01) regression equations, which are shown in Table 4.

Estimation of production by empirical relations

Table 5 gives the results of the comparison of production values calculated by the original authors for four different communities with values estimated by means of the regression of P on  $\bar{B}$  & W. At the population level, the deviations between calculated and estimated production reach several hundred per cent in some cases (mean = 74 %), whereas at the community level (i.e. sum of all species) the largest deviation is 66 % only (mean = 21 %).

Discussion

Many authors who analyzed the empirical relations between productivity and other parameters of animal populations made quite rigorous demands on the quality of the data to be included in their calculations. In contrast, I tried to maximize the number of data, although there is a considerable spread in the quality of the publications which are included here. This seems to be a more advantageous approach, if we assume that statistical error (i.e. the difference between the true values of B, W, P and P/ $\bar{B}$  and the published values) is randomly distributed among these data.

Up to now, the most common approach to establish empirical relations between productivity and other population parameters for macrobenthic invertebrates was the correlation of the P/ $\bar{B}$ -ratio to various parameters (see introduction), which are more or less equivalent to average body weight. All of these depend on the negative exponential relation of metabolic rate to body weight. This relation has been found here again, but it is not the main topic of this paper.

The results presented here show that P/ $\bar{B}$  does not depend on  $\bar{B}$  at an empirical basis. This is in contrast to several single species studies, which demonstrated density-dependent reductions in productivity or related parameters, see e.g. MORRISSEY (1987) or ZAJAK (1986). In an ecological context, the independence of P/ $\bar{B}$  from  $\bar{B}$  indicates that intra-

specific density-dependent mechanisms, e.g. intraspecific competition for food, do not cause the observed variability in the  $P/\bar{B}$ -ratio of various species and populations. The remaining variance of the  $P/\bar{B}$  data, which is not explained by  $W$ , must be due to other factors like interspecific competition, predation or environmental parameters which were not taken into account for this investigation.

Table 5. Comparison of calculated production values with values estimated from multiple regression equations

$P_{calc}$ : Production calculated by the original authors  
 $P_{est}$ : Computed with multiple regression of  $P$  on  $\bar{B}$  &  $W$ .  
 Original authors: Community No. 1: Buchanan & Warwick 1974  
 No. 2: Warwick & Price 1975  
 No. 3: Warwick et al. 1978  
 No. 4: Warwick & George 1980

Species	$\bar{B}$ gAFDW m <sup>-2</sup>	$W$ gAFDW	$P_{calc}$ gAFDW m <sup>-2</sup> y <sup>-1</sup>	$P_{est}$ gAFDW m <sup>-2</sup> y <sup>-1</sup>	Deviation %
Community No. 1					
<i>Abra nitida</i>	0.11	0.00688	0.12	0.11	-8
<i>Amnionrypane aulogaster</i>	0.17	0.00181	0.36	0.33	-8
<i>Chetozoeone setosa</i>	0.04	0.00048	0.05	0.09	60
<i>Heteromastus filiformis</i>	0.29	0.00129	0.30	0.59	97
<i>Lumbrineris fragilis</i>	0.06	0.00200	0.08	0.11	38
<i>Spiophanes kroeyeri</i>	0.12	0.00347	0.20	0.21	5
<i>Calocaris macandreae</i>	1.21	0.12100	0.14	0.64	286
Sum	2.00	-	1.25	2.08	66
Community No. 2					
<i>Macoma balthica</i>	0.34	0.00692	0.31	0.34	10
<i>Mya arenaria</i>	5.34	0.02828	2.66	4.10	54
<i>Scrobicularia plana</i>	2.15	0.03746	0.48	1.98	198
<i>Ampharete acutifrons</i>	0.43	0.00113	2.32	0.89	-62
<i>Nephtys hombergii</i>	3.95	0.01407	7.34	6.41	-13
Sum	12.41	-	13.11	13.17	<1
Community No. 3					
<i>Donax vittatus</i>	0.34	0.00449	0.72	0.39	-46
<i>Ensis siliqua</i>	5.10	0.82524	1.37	1.45	6
<i>Pharus legumen</i>	28.82	0.16707	16.12	13.36	-16
<i>Tellina fabula</i>	0.33	0.00162	0.29	0.50	72
<i>Venus striatula</i>	1.50	0.06608	0.62	0.84	35
<i>Glycera alba</i>	0.29	0.00380	0.28	0.52	86
<i>Magelona papillicornis</i>	0.63	0.00032	0.69	1.52	120
<i>Signation mathildae</i>	0.38	0.01667	0.17	0.57	235
<i>Spiophanes bombyx</i>	0.69	0.00061	3.35	1.55	-54
<i>Tharyx marioni</i>	0.02	0.00011	0.02	0.05	150
<i>Ophiura texturata</i>	0.67	0.06569	0.46	0.47	2
Sum	38.77	-	24.09	21.42	-11
Community No. 4					
<i>Abra alba</i>	0.30	0.00572	0.41	0.32	-22
<i>Nucula turrida</i>	2.62	0.00345	1.35	3.44	155
<i>Spisula elliptica</i>	6.25	0.00770	10.29	6.69	-35
<i>Ampharete firmarctica</i>	0.02	0.00159	0.03	0.04	33
<i>Nephtys hombergii</i>	0.46	0.00106	0.37	0.96	159
<i>Diastylis rathkei</i>	0.33	0.00223	0.40	0.71	78
<i>Ophiura texturata</i>	0.20	0.02048	0.10	0.19	90
Sum	10.16	-	12.95	12.35	-5

From the user's point of view the absent relation between  $P/\bar{B}$  and  $\bar{B}$  may be the main disadvantage of any empirical regression which uses  $P/\bar{B}$  as the dependent variable, because  $\bar{B}$  is required to estimate  $P$  from  $P/\bar{B}$  also, but  $\bar{B}$  is not used in this way to increase the precision of the estimate of  $P$ .

The multiple regression equations between  $P$  and  $\bar{B}$  &  $W$  are proposed here as a tool for estimating the annual production of macrobenthic communities. Table 5 shows that - as suggested by WARWICK (1980) - the empirical relations between  $P$  and  $\bar{B}$  &  $W$  result in reasonable estimates of community production, although there are large differences between calculated and estimated production at the species level. In three of the four examples presented, the deviations in the estimates of community production are below 20%. With respect to the first example (BUCHANAN and WARWICK 1974), the deviation is much higher (66%). This may be due to the extraordinary low standing stock and production of this community, which is a strong outlier with respect to the data base of the equations. A disadvantage of this comparison is the fact that the production data of these four communities are part of the data basis of the empirical regressions which are to be tested here. Only independent data on community production allow a real test on the validity of the regression equations. Those data, however, are not available up to now.

One important question is the establishment of confidence limits for an estimate of community production derived from these regression equations. If we estimate the production of a single population only, the 95% confidence limits of the predicted production value are the appropriate parameter, which can be calculated according to DRAPEL and SMITH (1981). The community production estimate consists of a sum of population production estimates, this makes the calculation of a statistically valid estimate of variability somewhat difficult.

To overcome this problem, I propose the application of "worst case" 95% limits to the estimate of community production. They simply consist of the sum of the lower and upper 95% confidence limits of the predicted production values of all populations. These "worst case" limits may be interpreted as the two opposite situations of systematic deviation, i.e. the  $P$ - $B$ - $W$  coordinates of all populations of the community in question show a deviation from the regression plane in the same direction. They are not real confidence limits, but they may be a useful measure of variation of the calculated community production, especially when different communities are to be compared. Table 6 contains all data required to calculate the 95% confidence limits of values estimated from the regression equations of  $P$  on  $\bar{B}$  and  $W$ .

In Table 7, the "worst case" 95% limits of the production estimates of the four communities of Table 5 are shown. In three of the four cases, these limits include the calculated production value; the calculated production of community 1, however, is below the lower 95% limit.

To sum up, the following procedure is proposed for the estimation of community production by means of the empirical relations presented here:

1. Calculate mean annual biomass  $\bar{B}$  (gAFDW) and mean individual weight  $W$  (gAFDW) of all populations which belong to the community in question.
2. Compute the estimate of annual production of each population by means of the regression equations of  $P$  on  $\bar{B}$  and  $W$  (Tables 3 and 4). Use the equations for taxonomic groups (Mollusca, Polychaeta, Crustacea) wherever possible.
3. Add up the population production values to get an estimate of community production.
4. Compute the "worst case" 95% limits of community production by adding up the 95% confidence limits of the predicted production values of the populations. Calculate these limits by means of the data given in Table 6 (e.g. according to DRAPEL and SMITH 1981).

In order to decrease the probability of extremely misleading estimates of annual production, the empirical relations presented here should not be applied to

Table 6. Computation of the 95% confidence limits of a production value which was predicted from the empirical regression equations. Notation according to DRAPER and SMITH (1981)

$$t = t_{n-2, 0.975} \text{ (one-tailed)}$$

$$s^2 = \text{Residual mean square of multiple regression}$$

$$X_0 = \begin{bmatrix} 1 \\ \log(\bar{B}_1) \\ \log(\bar{W}_1) \end{bmatrix}$$

$$X = \begin{bmatrix} 1 & \log(\bar{B}_1) & \log(\bar{W}_1) \\ \vdots & \vdots & \vdots \\ 1 & \log(\bar{B}_n) & \log(\bar{W}_n) \end{bmatrix}$$

$$C = (X'X)^{-1}$$

$$\text{Calculation: lower 95\% limit} = \log(P_1) - t * \sqrt{s^2 * X_0' C X_0}$$

$$\text{upper 95\% limit} = \log(P_1) + t * \sqrt{s^2 * X_0' C X_0}$$

Taxon	S <sup>2</sup>	C		
Mollusca	0.127	0.037222	-0.012411	0.014153
		-0.012411	0.008481	-0.004876
Polychaeta	0.091	0.014153	-0.004876	0.006434
		-0.004876	0.038806	0.044492
Crustacea	0.046	0.122647	0.036950	-0.016493
		-0.038806	0.036950	0.018218
Total Data	0.118	0.044492	-0.016493	0.160610
		0.515868	-0.042474	-0.017372
		-0.042474	0.035476	0.051981
		0.160610	-0.017372	-0.001241
		-0.003395	0.001047	0.000114
		0.001047	0.002847	0.000004
		-0.000114	0.000114	

Table 7. The effect of the 95% confidence intervals of the multiple regression functions on the estimate of community production

P<sub>calc</sub>: Production calculated by the original authors.  
P<sub>est</sub>: Computed with multiple regression of P on  $\bar{B}$  &  $\bar{W}$ .  
"worst case" 95% limits: Sum of the lower and upper 95% confidence limits of the predicted production values

Community No.	P <sub>calc</sub>	P <sub>est</sub>	"worst case" 95% limits
			lower upper
1	1.25	2.08	1.51 2.87
2	13.11	13.17	11.08 15.70
3	24.09	21.42	17.82 25.84
4	12.95	12.35	10.63 14.35

1. communities dominated by one species with extraordinary high biomass like mussel beds,
  2. communities with only a few (<5) species,
  3. communities exposed to very stressing environmental conditions.
- If situations of this kind are excluded, the empirical relations between P and  $\bar{B}$  &  $\bar{W}$  presented here seem to be useful for a first rough estimate of production of shallow water (down to 300 m water depth) macrobenthic communities. However, the existence of empirical relations of this kind should not keep students and scientists from carrying out production studies in the "good old way". The results presented here show very clearly that production estimates based on empirical relations are close to reality up to the average, but they may be quite misleading in certain cases.

Geographical latitude and water depth may be two abiotic parameters which influence

benthic secondary production and P/ $\bar{B}$ -ratio via primary production, water temperature and amount of available food. However, with respect to these parameters the data do not cover a sufficient range for a valid analysis (Fig. 1 and Fig. 2). Further work on macrobenthic secondary production in the neglected areas may provide additional data to overcome this problem.

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